



Probabilistic deconvolution of the distribution of relaxation times from multiple electrochemical impedance spectra

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HIGHLIGHTS

- Enhances multi-spectra analysis using quasi-Gaussian process.
- Evaluates DRT inversion's credibility.
- Accurately recovers DRT from noisy data.
- Provides a unified probabilistic approach for DRT analysis.

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ABSTRACT

Electrochemical impedance spectroscopy (EIS) is widely used to study the properties of electrochemical materials and systems. However, analyzing EIS data remains challenging. Among various analysis methods, the distribution of relaxation times (DRT) has emerged as a novel non-parametric approach capable of providing timescale information. Among the various DRT inversion methods, those based on Gaussian processes (GP) are particularly promising because they provide uncertainty estimates for both EIS and DRT. However, current GP-based DRT implementations can only handle one spectrum at a time. This work extends these models to allow concurrent analysis of multiple impedance spectra as a function of experimental conditions. The new method, called the quasi-Gaussian process distribution of relaxation times, treats the DRT as a GP with respect to the experimental state and as a finite approximation of a positively constrained GP with respect to timescales. This new DRT inversion approach is validated against noise-corrupted artificial EIS data and applied to experimental data, allowing us to expedite EIS data analysis of multiple EIS data from a probabilistic perspective.

1. Introduction

Electrochemical impedance spectroscopy (EIS) is a versatile characterization and diagnostic technique widely used in electrochemistry applications [1–4] to study batteries [5–8], fuel cells [7,9,10], solar cells [11,12], electrolyzers [13–15], and beyond [16–21]. The distribution of relaxation times (DRT) has recently emerged as a promising non-parametric approach to analyze impedance spectra [22,23] and has been applied to identify electrode processes [24,25], estimate battery state-of-health [26–29], and assess the performance of fuel cells [30,31].

Furthermore, the DRT has been used to develop equivalent circuit models without requiring prior circuit architecture information [32–35]. Assuming that the experimental impedance depends on both frequency, f , and an experimental state, ψ ,¹ the DRT impedance model, $Z_{\text{DRT}}(f, \psi)$, can be expressed in terms of f and ψ as shown:

$$Z_{\text{DRT}}(f, \psi) = i2\pi f L_0(\psi) + R_\infty(\psi) + \int_{-\infty}^{+\infty} \frac{\gamma(\log \tau, \psi)}{1 + 2\pi i f \tau} d \log \tau \quad (1)$$

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¹ The experimental state, ψ , is encoded in a vector. For example, for a battery experiment, the vector state could comprise the state of charge (SoC) and cycle number (n), implying $\psi = (\text{SoC} \ n)^\top$. While the theory in this article is developed with dependence on an experimental state vector, for the sake of clarity and without loss of generality, the DRT inversion from both synthetic and real experimental data is shown with respect to ψ , a scalar experimental state.

where $L_0(\boldsymbol{\psi})$, $R_\infty(\boldsymbol{\psi})$, and $\gamma(\log \tau, \boldsymbol{\psi})$ are the inductance, resistance, and the DRT, respectively. It should be noted that, to account for the dependence on the experimental state, we have encoded the $L_0(\boldsymbol{\psi})$, $R_\infty(\boldsymbol{\psi})$, and $\gamma(\log \tau, \boldsymbol{\psi})$ as functions parameterized on the experimental state ($\boldsymbol{\psi}$). In particular, $\gamma(\log \tau, \boldsymbol{\psi})$ is explicitly written as dependent on the log-timescale ($\log \tau$) with a parameterization on $\boldsymbol{\psi}$. This is consistent with our previous articles [22,36–38], where the DRT problem does not change relative to its formulation [37,38], but rather, the parametrization is shown explicitly.

Deconvolving $\gamma(\log \tau, \boldsymbol{\psi})$ from EIS data is challenging, as it requires solving an ill-posed inverse problem [39,40]. Regularized regression [41–46], deep neural networks [37], and other methods [33,34,47–54] have been developed for this purpose. Our group introduced a probabilistic approach for DRT deconvolution, reformulating regularized regression in a Bayesian context [55], including the Gaussian-process-based DRT (GP-DRT) model [56]. However, the GP-DRT method can produce negative DRT values (if a non-negativity constraint is not imposed through some inducing points) and is limited to the imaginary part of the EIS spectrum. To address these challenges, we developed the finite GP-DRT (fGP-DRT) framework [36], which inherits the key properties of GP-DRT, utilizes both components of the impedance, and enforces a non-negativity constraint on the DRT. While promising, the current fGP-DRT's limitation in deconvolving one EIS spectrum at a time restricts the analysis to a single experimental state without considering how several EIS spectra are correlated. This limitation could hinder the identification of electrochemical processes that evolve across conditions, such as oxygen partial pressure or temperature in fuel cells or cycle number, state of charge, or health in batteries.

This article addresses this limitation by extending the fGP model to explicitly incorporate dependence on the experimental state, $\boldsymbol{\psi}$. Specifically, $\gamma(\log \tau, \boldsymbol{\psi})$ is modeled as a GP with respect to $\boldsymbol{\psi}$ and a finite GP (fGP) with respect to $\log \tau$. Additionally, both $L_0(\boldsymbol{\psi})$ and $R_\infty(\boldsymbol{\psi})$ are assumed to be GPs. This framework is termed quasi-GP DRT (qGP-DRT) as it combines a GP (with respect to $\boldsymbol{\psi}$) with a fGP (with respect to $\log \tau$). This advancement allows for the simultaneous analysis of multiple EIS spectra (Fig. 1), providing a unified probabilistic DRT inversion framework across experimental conditions. To ensure its consistency, the qGP-DRT model was tested against noise-corrupted EIS synthetic experiments, showing that the deconvolved DRT closely matched analytical results. Furthermore, validation against real experimental data from fuel cells and batteries showed consistent DRT inversion against equivalent circuit models. Importantly, the qGP-DRT framework provides Bayesian probabilistic credible bands of impedance and DRT, allowing the assessment of the uncertainty with respect to both. In short, this work introduces a novel probabilistic approach for EIS data analysis based on Gaussian processes, paving the way for new research avenues in DRT analysis of multiple EIS spectra at a time.

2. Methods

As outlined in the Introduction, the qGP-DRT framework enhances the fGP-DRT model to handle multiple impedance spectra, allowing multidimensional experimental-dependent DRT inversion. This section covers its theoretical background, sampling process, and scoring used to check the consistency of the qGP-DRT model.

2.1. Theory

The qGP-DRT model assumes that $\gamma(\log \tau, \boldsymbol{\psi})$ is a zero-mean GP with respect to the experimental state vector, $\boldsymbol{\psi}$, and a fGP with respect to $\log \tau$ [36].² Additionally, the model approximates $\gamma(\log \tau, \boldsymbol{\psi})$ as

² We may write loosely that $\gamma(\boldsymbol{\psi}|\log \tau) \sim \mathcal{GP}(0, k(\boldsymbol{\psi}, \boldsymbol{\psi}'))$ with kernel $k(\boldsymbol{\psi}, \boldsymbol{\psi}')$, where $\gamma(\boldsymbol{\psi}|\log \tau)$ indicates that $\log \tau$ is fixed.

$$\gamma(\log \tau, \boldsymbol{\psi}) \approx \sum_{n=1}^N \gamma(\log \tau_n, \boldsymbol{\psi}) \phi_n(\log \tau) \quad (2)$$

where $\phi_n(\log \tau)$ is a basis function [36], $\gamma(\log \tau_n, \boldsymbol{\psi})$ is the DRT at the nodal points $\log \tau_n$ and N is the number of collocation point with $\log \tau = (\log \tau_1, \log \tau_2, \dots, \log \tau_N)^\top$.

To illustrate the approach developed, we consider first only EIS data from a pair of spectra at two experimental states, $\boldsymbol{\psi}_p$ and $\boldsymbol{\psi}_q$. We denote the corresponding data as \mathbf{Z}_p and \mathbf{Z}_q , which are the vectors concatenating the real and imaginary components of p^{th} and q^{th} experiments, respectively [36]. The vectors of probed frequencies can be generically denoted as $\mathbf{f}_p = (f_{p,1}, f_{p,2}, \dots, f_{p,M_p})^\top$ and $\mathbf{f}_q = (f_{q,1}, f_{q,2}, \dots, f_{q,M_q})^\top$, where M_p and M_q are not necessarily identical positive integers.

We consider the DRT discretization outlined by equation (2), such that the discretization vector for the p^{th} experiment is $\mathbf{x}_p = (L_0(\boldsymbol{\psi}_p), R_\infty(\boldsymbol{\psi}_p), \gamma(\log \tau, \boldsymbol{\psi}_p)^\top)^\top$. Then, we can rewrite (1) as [36]

$$\mathbf{Z}_p = \mathbf{A}_p \mathbf{x}_p + \boldsymbol{\varepsilon}_p \quad (3a)$$

$$\mathbf{Z}_q = \mathbf{A}_q \mathbf{x}_q + \boldsymbol{\varepsilon}_q \quad (3b)$$

where \mathbf{A}_p and \mathbf{A}_q are obtained by discretizing the real and imaginary parts of the DRT impedances using the real ($\mathbf{A}_{\text{re},p}$) and imaginary ($\mathbf{A}_{\text{im},p}$), and the real ($\mathbf{A}_{\text{re},q}$) and imaginary ($\mathbf{A}_{\text{im},q}$), component matrices, respectively, so that \mathbf{A}_p is given by [36]

$$\mathbf{A}_p = \begin{pmatrix} \mathbf{0} & \mathbf{1} & \mathbf{A}_{\text{re},p} \\ 2\pi \mathbf{f}_p & \mathbf{0} & \mathbf{A}_{\text{im},p} \end{pmatrix} \quad (4)$$

and $\boldsymbol{\varepsilon}_p \sim \mathcal{N}(0, \sigma_n^2 \mathbf{I}_M)$ where σ_n is a scalar and \mathbf{I}_M is $M \times M$ identity matrix. For the q^{th} experiment, definitions are analogous.

Assuming, the resistance, $R_\infty(\boldsymbol{\psi})$, and inductance, $L_0(\boldsymbol{\psi})$, are zero-mean GPs, such that $R_\infty(\boldsymbol{\psi}) \sim \mathcal{GP}(0, k_R(\boldsymbol{\psi}, \boldsymbol{\psi}'))$ and $L_0(\boldsymbol{\psi}) \sim \mathcal{GP}(0, k_L(\boldsymbol{\psi}, \boldsymbol{\psi}'))$ with kernels $k_R(\boldsymbol{\psi}, \boldsymbol{\psi}')$ and $k_L(\boldsymbol{\psi}, \boldsymbol{\psi}')$, respectively.

It follows from (3) that

$$\begin{pmatrix} \mathbf{x}_p \\ \mathbf{x}_q \end{pmatrix} \sim \mathcal{N}(0, \boldsymbol{\Gamma}) \quad (5)$$

where the covariance matrix, $\boldsymbol{\Gamma}$, is defined as

$$\boldsymbol{\Gamma} = \begin{pmatrix} \boldsymbol{\Gamma}_{pp} & \boldsymbol{\Gamma}_{pq} \\ \boldsymbol{\Gamma}_{pq}^\top & \boldsymbol{\Gamma}_{qq} \end{pmatrix} \quad (6)$$

and diagonal and off-diagonal blocks ($\boldsymbol{\Gamma}_{pp}$ and $\boldsymbol{\Gamma}_{pq}$, respectively) are

$$\boldsymbol{\Gamma}_{pp} = \begin{bmatrix} k_R(\boldsymbol{\psi}_p, \boldsymbol{\psi}_p) & 0 & 0 \\ 0 & k_L(\boldsymbol{\psi}_p, \boldsymbol{\psi}_p) & 0 \\ 0 & 0 & \mathbf{K}_{pp} \end{bmatrix} \quad (7a)$$

$$\boldsymbol{\Gamma}_{pq} = \begin{bmatrix} k_R(\boldsymbol{\psi}_p, \boldsymbol{\psi}_q) & 0 & 0 \\ 0 & k_L(\boldsymbol{\psi}_p, \boldsymbol{\psi}_q) & 0 \\ 0 & 0 & \mathbf{K}_{pq} \end{bmatrix} \quad (7b)$$

We can explicitly write the diagonal and off-diagonal blocks as

$$(\mathbf{K}_{pp})_{nm} = k_\gamma((\log \tau_n, \boldsymbol{\psi}_p), (\log \tau_m, \boldsymbol{\psi}_p)) \quad (8a)$$

$$(\mathbf{K}_{pq})_{nm} = k_\gamma((\log \tau_n, \boldsymbol{\psi}_p), (\log \tau_m, \boldsymbol{\psi}_q)) \quad (8b)$$

where $n, m = 1, 2, \dots, N$. We can now generalize to the case of $N_\boldsymbol{\psi}$

experimental conditions so that $\mathbf{x} = \begin{pmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \\ \vdots \\ \mathbf{x}_{N_\boldsymbol{\psi}} \end{pmatrix}$ and $\mathbf{Z}_{\text{exp}} = \begin{pmatrix} \mathbf{Z}_1 \\ \mathbf{Z}_2 \\ \vdots \\ \mathbf{Z}_{N_\boldsymbol{\psi}} \end{pmatrix}$ are

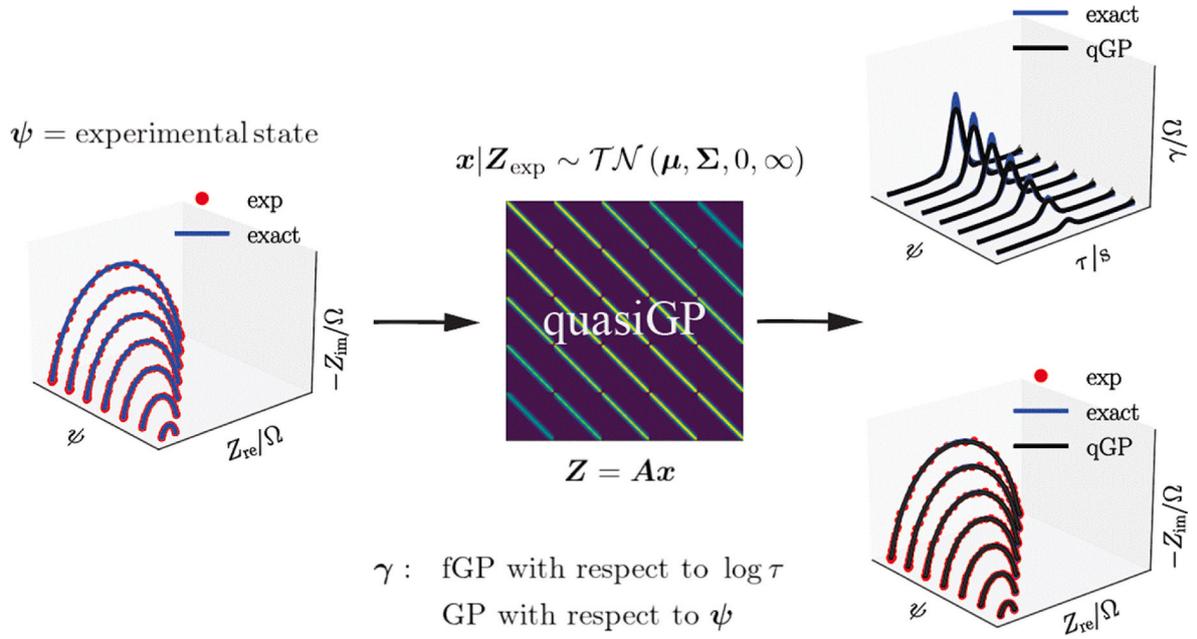


Fig. 1. Schematic illustration of the qGP-DRT framework.

the vectors of discretized DRTs and experimental impedances, respectively. Each vector \mathbf{x}_k and \mathbf{Z}_k (for $k = 1, 2, \dots, N_\psi$) can be defined in an identical way as above. As in the fGP model [36], the distribution of \mathbf{x} conditioned to the experimental data \mathbf{Z}_{exp} is a truncated multivariate normal such that [36,56]

$$\mathbf{x}|\mathbf{Z}_{\text{exp}} \sim \mathcal{TN}(\boldsymbol{\mu}, \boldsymbol{\Sigma}, \mathbf{0}, \infty) \quad (9)$$

where the mean, $\boldsymbol{\mu}$, and covariance matrix, $\boldsymbol{\Sigma}$, are defined as

$$\boldsymbol{\mu} = \text{abs}(\boldsymbol{\Gamma}\mathbf{A}^\top(\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top + \sigma_n^2 \mathbf{I}_{N_\psi})^{-1}\mathbf{Z}_{\text{exp}}) \quad (10a)$$

$$\boldsymbol{\Sigma} = \boldsymbol{\Gamma} - \boldsymbol{\Gamma}\mathbf{A}^\top(\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top + \sigma_n^2 \mathbf{I}_{N_\psi})^{-1}\mathbf{A}\boldsymbol{\Gamma} \quad (10b)$$

and $\text{abs}(\cdot)$ indicates elementwise absolute value, \mathbf{I}_{N_ψ} is the $N_\psi \times N_\psi$ identity matrix, \mathbf{A} is a block-diagonal matrix such that $\mathbf{A} = \text{diag}(\mathbf{A}_1, \mathbf{A}_2, \dots, \mathbf{A}_{N_\psi})$, and finally $\boldsymbol{\Gamma}$ is generalized relative to (6) as

$$\boldsymbol{\Gamma} = \begin{pmatrix} \boldsymbol{\Gamma}_{11} & \boldsymbol{\Gamma}_{12} & \cdots & \boldsymbol{\Gamma}_{1N_\psi} \\ \boldsymbol{\Gamma}_{12}^\top & \boldsymbol{\Gamma}_{22} & \ddots & \boldsymbol{\Gamma}_{2N_\psi} \\ \vdots & \vdots & \ddots & \vdots \\ \boldsymbol{\Gamma}_{1N_\psi}^\top & \boldsymbol{\Gamma}_{2N_\psi}^\top & \cdots & \boldsymbol{\Gamma}_{N_\psi N_\psi} \end{pmatrix} \quad (11)$$

where the definition of each block is analogous to the block definition given in (6).

2.1.1. Kernels

To compute the kernel blocks in (11) using (7) and (8), we employ the following kernels [36,56]:

$$k_f((\log \tau, \boldsymbol{\psi}), (\log \tau', \boldsymbol{\psi}')) = \sigma_v^2 \exp\left(-\frac{(\log \tau - \log \tau')^2}{2\ell_f^2}\right) \exp\left(-\frac{\|\boldsymbol{\psi} - \boldsymbol{\psi}'\|_2^2}{2\ell_\psi^2}\right) \quad (12a)$$

$$k_R(\boldsymbol{\psi}, \boldsymbol{\psi}') = \sigma_R^2 \exp\left(-\frac{\|\boldsymbol{\psi} - \boldsymbol{\psi}'\|_2^2}{2\ell_R^2}\right) \quad (12b)$$

$$k_L(\boldsymbol{\psi}, \boldsymbol{\psi}') = \sigma_L^2 \exp\left(-\frac{\|\boldsymbol{\psi} - \boldsymbol{\psi}'\|_2^2}{2\ell_L^2}\right) \quad (12c)$$

where, σ_v^2 , σ_R^2 , σ_L^2 , are the kernel's prefactors, ℓ_f , ℓ_R , ℓ_L , ℓ_ψ denote length scales, and $\|\cdot\|_2$ is the Euclidean norm.

2.1.2. Hyperparameter optimization

Unless explicitly stated, the hyperparameter vector $\boldsymbol{\theta} = (\sigma_n, \sigma_v, \sigma_R, \sigma_L, \ell_f, \ell_\psi, \ell_R, \ell_L)^\top$ was obtained by maximizing the unbounded evidence, i.e., $p(\mathbf{Z}_{\text{exp}}|\boldsymbol{\theta})$, which is given by

$$\log(p(\mathbf{Z}_{\text{exp}}|\boldsymbol{\theta})) = \frac{1}{2} \mathbf{Z}_{\text{exp}}^\top (\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top + \sigma_n^2 \mathbf{I}_{N_\psi})^{-1} \mathbf{Z}_{\text{exp}} - \frac{1}{2} \log(|\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top + \sigma_n^2 \mathbf{I}_{N_\psi}|) - \frac{(M_{\text{tot}} + 2)}{2} \log(2\pi) \quad (13)$$

where $M_{\text{tot}} = \sum_{k=1}^{N_{\text{exp}}} M_k$ with M_k being the number of frequencies probed during the k^{th} experiment.

As calculating the direct inverse and determinant of the large matrix in (13) can be numerically challenging, $\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top$ was approximated using a truncated singular value decomposition, such that $\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top \approx \mathbf{U}_r \mathbf{S}_r \mathbf{V}_r^\top$, where matrices \mathbf{U}_r and \mathbf{V}_r contain the first r columns and rows of the orthogonal matrices \mathbf{U} and \mathbf{V} , respectively, and $\mathbf{S}_r = \text{diag}(\sigma_1, \sigma_2, \dots, \sigma_r)$ contains the r largest singular values of $\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top$ with $\sigma_1 > \sigma_2 > \dots, \sigma_r$ (panel (a) of Fig. S1). Then, $(\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top + \sigma_n^2 \mathbf{I})^{-1}$ was approximated using the Sherman-Morrison-Woodbury's formula [57] as

$$(\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top + \sigma_n^2 \mathbf{I})^{-1} \approx \frac{1}{\sigma_n^2} \mathbf{I} - \frac{1}{\sigma_n^4} \mathbf{U}_r \left(\text{diag} \left(\frac{\sigma_1 \sigma_n^2}{\sigma_1 + \sigma_n^2}, \frac{\sigma_2 \sigma_n^2}{\sigma_2 + \sigma_n^2}, \dots, \frac{\sigma_r \sigma_n^2}{\sigma_r + \sigma_n^2} \right) \right) \mathbf{V}_r^\top \quad (14)$$

and the determinant $|\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top + \sigma_n^2 \mathbf{I}|$ was approximated using the matrix inversion lemma [58] as

$$|\mathbf{A}\boldsymbol{\Gamma}\mathbf{A}^\top + \sigma_n^2 \mathbf{I}| \approx \prod_{k=1}^n (\sigma_n^2 + \sigma_k) \quad (15)$$

2.1.3. Sampling

Relative to the previously used Hamiltonian Monte Carlo sampler [36], we added an adaptive step size [59] to sample \mathbf{x} from (9). The adaptive step size optimizes exploration in Hamiltonian dynamics by monitoring acceptance rates: small steps slow sampling, and large steps introduce inaccuracy. The algorithm adjusts step size to balance efficiency (high acceptance) and exploration (finding diverse states). The step size dynamically adjusts based on the acceptance rate over 50-sample intervals. Rates below 90 % trigger a step-size decrease for more careful exploration, while rates above 90 % lead to an increase of the step size for faster exploration. This ensures efficient sampling within the desired constraints.

During sampling, 10,000 samples were generated with the initial 1000 samples being discarded as burn-in. Next, having generated 9000 samples of \mathbf{x} , the corresponding impedances, \mathbf{Z} , were obtained as $\mathbf{Z} = \mathbf{A}\mathbf{x}$. The 1st to 99th percentile credible intervals around the means were visualized as gray regions in corresponding plots.

2.2. Scoring the quality of the DRT and EIS recovery

To evaluate the quality of DRT and impedance recoveries, the following normalized mean square errors were used:

$$\text{MSE}_{\text{norm},\gamma} = \frac{1}{N_\psi} \sum_{k=1}^{N_\psi} \mathbb{E} \left(\frac{\|\gamma_{\text{exact},k} - \gamma_k\|^2}{\|\gamma_{\text{exact},k}\|^2} \right) \quad (16a)$$

$$\text{MSE}_{\text{norm},\mathbf{Z}} = \frac{1}{N_\psi} \sum_{k=1}^{N_\psi} \mathbb{E} \left(\frac{\|\mathbf{Z}_{\text{exact},k} - \mathbf{Z}_k\|^2}{\|\mathbf{Z}_{\text{exact},k}\|^2} \right) \quad (16b)$$

where $\mathbb{E}[\cdot]$ denotes expected value computed from the samples obtained as described in the previous section, $\gamma_{\text{exact},k}$ is the exact DRT vector corresponding to exact impedance vector $\mathbf{Z}_{\text{exact},k}$, γ_k and $\mathbf{Z}_{\text{exact},k}$ are the correspondingly recovered DRT and impedance vectors, respectively. As above, γ is discretized with respect to $\log \tau$.

2.3. Artificial experiment generation

The noise-corrupted artificial data were generated in the $10^{-2} - 10^6$ Hz frequency range with 10 points per decade by adding noise to the exact impedance $\mathbf{Z}_{\text{exact}}$. Unless explicitly stated, the error standard deviation was set to $\sigma_n^{\text{exp}} = 0.5 \Omega$.³

3. Results

Synthetic experiments were first used to validate the consistency and robustness of the qGP-DRT model. Then, the model was tested against real data from a fuel cell and battery under diverse experimental conditions.

3.1. Artificial experiments

We first considered EIS data originating from single ZARC, Havriliak-Negami, $2 \times$ ZARC, Gerischer, and piecewise constant models [36,56,60] (their analytical expressions are reported elsewhere [41]). Without loss of generality, the experimental state is taken to be ψ , a scalar.

3.1.1. Single ZARC and Havriliak-Negami models

We initially investigated a single ZARC model consisting of an ohmic resistor (R_∞) in series with a parallel circuit containing a resistor (R_{ct}) and a constant phase element (CPE) [37]. Six synthetic experiments

³ σ_n^{exp} is the standard deviation of the noise in the artificial experiments analogous to $\epsilon \sim \mathcal{N}(0, \sigma_n^{\text{exp}} \mathbf{I})$ in equation (3). Instead, σ_n is one of the hyperparameters of the qGP-DRT model originating from equation (3).

were generated with $R_{\text{ct}}(\psi)$ exhibiting a dependence on ψ . Fig. 2 depicts the median DRTs and impedances (panels (a) and (c)) recovered with the qGP-DRT model for $\psi = 0.0, 0.2, \dots, 1.0$. All DRTs and impedances (panels (a), (c) & (e)) were recovered accurately, as evidenced by the low $\text{MSE}_{\text{norm},\gamma} = 1.73 \times 10^{-2}$ and $\text{MSE}_{\text{norm},\mathbf{Z}} = 9.89 \times 10^{-6}$ values (Table S7). DRT peaks were consistently identified for all experimental conditions, with a noticeable peak decrease as ψ increased (Fig. 2 (e)). The recovered qGP-DRT impedances closely matched the corresponding exact impedances (panel (a) of Fig. 2). To evaluate the computational scaling of the qGP-DRT framework, we generated 5 synthetic experiments with each $N_\psi = 10, 20, 30$, and 40, and recorded the computational time.⁴ We noted the scalability of the qGP-DRT model, given that 20 experiments took approximately 17 min (Figure S1 (b)).

We then investigated a scenario where $\tau_0(\psi)$ exhibited a dependence on ψ . Eleven EIS spectra were subsequently generated for $\psi = 0.0, 0.2, \dots, 1.0$. Panels (a), (b), and (c) of Fig. 3 show the magnitudes, qGP DRTs, and 2D qGP-DRT plots, respectively. Notably, the recovered qGP DRTs demonstrated characteristic response shifts as ψ varied (panels (c) and (d) of Fig. 3). Furthermore, we applied the qGP-DRT model to analyze multiple EIS spectra derived from the Havriliak-Negami model. As for the single ZARC case, the DRTs and impedances were well recovered (see panels (b), (d), and (f) of Fig. 2).

3.1.1.1. Influence of the experimental noise. We evaluated the influence of experimental noise on the qGP-DRT framework. We employed a single ZARC model and generated synthetic experiments for various noise levels ($\sigma_n^{\text{exp}} = 1, 1.5$, and 2.0Ω) and $\psi = 0.0, 0.2, \dots, 1.0$. As above, we obtained hyperparameters optimally by maximizing the evidence at each noise level. The recovered qGP-DRTs and EIS are presented in Figs. S2 and S3. Even at elevated noise levels (1.5Ω and 2.0Ω), the recovered DRTs and impedances closely matched their corresponding exact values (see Fig. S3).

Additionally, we considered frequency-dependent noise models defined elsewhere [41], with the artificial EIS data being generated at noise level $\sigma_n^{\text{exp}} = 0.02\Omega$. Consistent with previous findings, the qGP DRTs and impedances closely matched the corresponding exact and experimental impedances (Fig. S4), as confirmed by the low values of $\text{MSE}_{\text{norm},\gamma}$ and $\text{MSE}_{\text{norm},\mathbf{Z}}$ for all σ_n^{exp} listed in Table S7.

3.1.1.2. Multiple ZARCs

We first studied the $2 \times$ ZARC model, which consists of two ZARCs (Section 2.1.1) in series. These circuits were chosen to exhibit overlapping, separated, or distant timescales (see the circuit parameters in Tables S2 and S3). Six synthetic EIS spectra were generated with $R_{\text{ct},2}(\psi)$ and $\tau_2(\psi)$ exhibiting a dependence on ψ .

For the case of the separated $2 \times$ ZARC ($\tau_1 = 1.0 \times 10^{-1}\text{s}$, $\tau_2 = 1.0 \times 10^{-3}\text{s}$), the Nyquist plots and the qGP DRTs are presented in panels (a) and (c) of Fig. S5, respectively. Consistent with our observations for the single ZARC model (Section 2.1.1), the qGP-DRT model successfully identified the DRT peaks and recovered the impedances across all experimental conditions. As shown in panels (c) and (e) of Fig. S5, the qGP DRTs exhibited two distinct peaks: one prominent for all ψ , and a second peak that diminishes with changing ψ . The overlapping $2 \times$ ZARC model (with $\tau_1 = 1.0 \times 10^{-1}\text{s}$, $\tau_2 = 1.0 \times 10^{-2}\text{s}$) was also well recovered, as demonstrated in panels (b) and (d) of Fig. S5.

These findings were further confirmed for the distant $2 \times$ ZARC and $3 \times$ ZARC models with the estimated median DRTs and impedances closely matching their corresponding exact values (see Figs. 4 and 5). Additionally, the low $\text{MSE}_{\text{norm},\gamma}$ and $\text{MSE}_{\text{norm},\mathbf{Z}}$ values in Table S7 corroborated the quality of the DRT and impedance recovery by the qGP-DRT model.

⁴ The clock time was measured using an Intel(R) Core (TM) i5-10300H CPU with a 2.50 GHz clock rate in a laptop computer equipped with 16 GB of RAM.

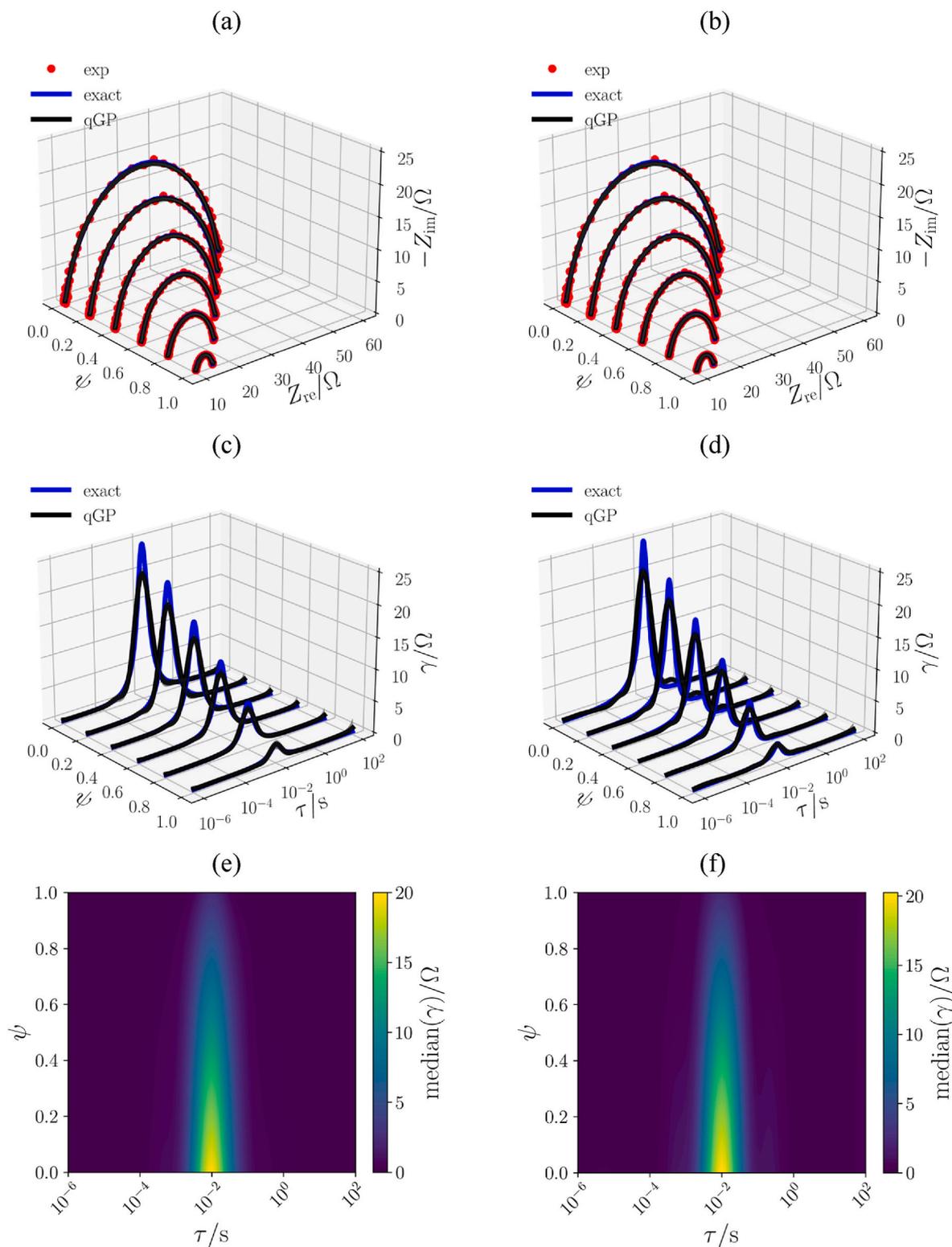


Fig. 2. Nyquist plots of the single ZARC (a) and Havriliak-Negami models (b) with $R_{ct}(\psi)$ exhibiting a dependence on ψ ($\sigma_n^{exp} = 0.5\Omega$). Panels (c) and (d) show the corresponding recovered qGP DRTs, with associated 2D contour plots of median DRTs in panels (e) and (f).

3.1.3. Discontinuous DRT

We extended the analysis to two discontinuous models, namely the Gerischer and piecewise constant models, whose circuit's parameters values are provided in Table S4. The corresponding recovered qGP DRTs and impedances are presented in Fig. S6. We observed that the qGP-DRT model approximately captured discontinuities.

3.2. Real experiments

Having established that the qGP-DRT model is consistent against well-controlled, analytical models, we evaluated its effectiveness in handling real-world data leveraging EIS spectra from a fuel cell and battery.

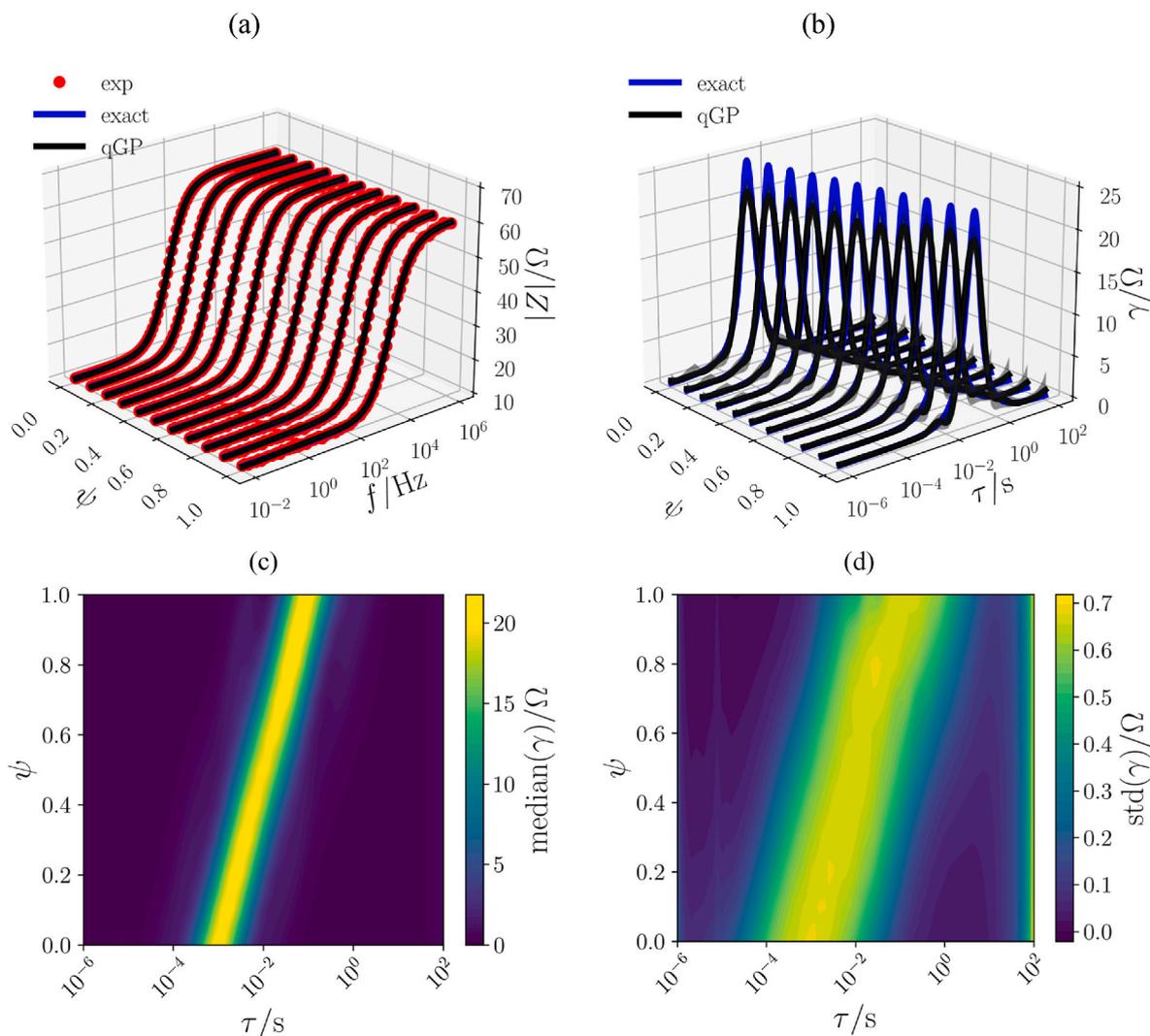


Fig. 3. (a) Magnitudes for the single ZARC model with $\tau_0(\psi)$ exhibiting dependence on ψ ($\sigma_n^{\text{exp}} = 0.5\Omega$). Panel (b) shows corresponding qGP DRTs, while panels (c) and (d) display 2D contour plots of the median and standard deviation, respectively.

3.2.1. Symmetric protonic ceramic fuel cell

We analyzed data from a symmetric cell featuring $\text{Ba}_{0.95}\text{La}_{0.05}\text{FeO}_{3-\delta}$ (BLF) as the electrode materials and a samarium-doped ceria as the electrolyte. The cell was operated at 550°C in an N_2/O_2 atmosphere under varying oxygen partial pressures ($p\text{O}_2 = 21, 40, 60, 80, \text{ and } 100\%$). EIS data were collected across the frequency range 100 mHz to 20 kHz with five points recorded per decade. The data were regressed using $2 \times \text{ZARC}$ ECM, with the parameters provided in Table S5. The qGP-DRT model successfully identified the DRT peaks for all experimental conditions, as shown in Fig. 6 (panels (c) and (e)). Additionally, the qGP-DRT model matched with the DRTtools,⁵ the fitted ECM and the experimental data (panel (a) and (c) Fig. 6), as evidenced by the low $\text{MSE}_{\text{norm},\gamma}$ and $\text{MSE}_{\text{norm},Z}$ values in Table S8.

3.2.2. Lithium-metal battery

Next, we investigated a lithium-ion battery with a LiFePO_4 (LFP) cathode, a lithium-metal anode, and 1M LiPF_6 in ethylene carbonate: diethyl carbonate (1:1 v/v) as an electrolyte [62]. EIS measurements

⁵ For the pyDRTtools settings, Gaussian radial basis functions for the discretization, and the full impedance spectrum were used. A regularization second-order derivative matrix was employed, and the regularization parameter was computed using the generalized cross-validation [41,61].

were collected at frequencies ranging from 0.1 Hz to 7 MHz with a 5C charge-discharge rate at cycles 30, 60, 90, and 120. The data were regressed using a $2 \times \text{ZARC}$ ECM, with the parameters provided in Table S6.

The qGP-DRT model successfully recovered the DRTs and impedances, the corresponding DRTs were in good agreement with DRTtools and ECM results (Fig. 6, panels (b) and (d)). Well-defined peaks were consistently identified $\tau = 10^{-4}$ s for all experimental conditions (Fig. 6 (f)). Furthermore, the low $\text{MSE}_{\text{norm},\gamma}$ and $\text{MSE}_{\text{norm},Z}$ values in Table S8 confirmed the effectiveness of the qGP-DRT regression. Additionally, the narrow credible bands in Fig. 6 (d) indicated low uncertainty and high confidence in the recovered DRTs.

4. Remarks and future directions

The qGP-DRT framework developed in this work enables simultaneous probabilistic analysis of multiple EIS spectra. We validated its effectiveness and consistency against both artificial and real datasets, with model hyperparameters optimally selected through evidence maximization. Despite its promise, several avenues for future improvements can be identified.

To enhance computational efficiency in EIS data analysis, we employed singular value decomposition for low-rank matrix approximation and adaptive step-size HMC sampling. However, challenges

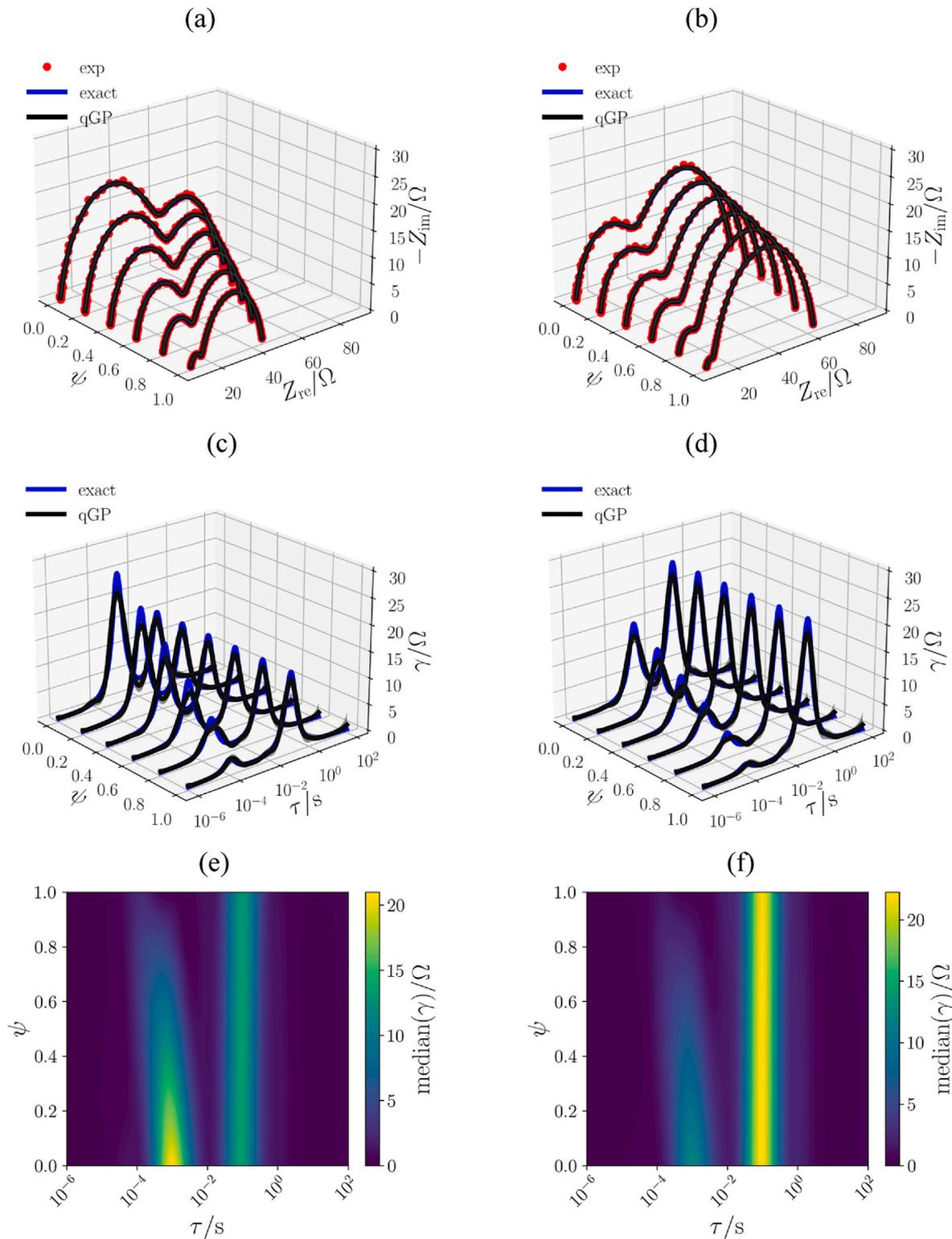


Fig. 4. Nyquist plots of the distant $2 \times \text{ZARC}$ model for two conditions: (a) $R_{ct,1} < R_{ct,2}$ and (b) $R_{ct,1} > R_{ct,2}$. In both conditions, $R_{ct,1}(\psi)$ and $\tau_2(\psi)$ exhibit a dependence on ψ ($\sigma_n^{\text{exp}} = 0.5\Omega$, $\tau_1 = 1.0 \times 10^{-1}$ s, and $\tau_2 = 1.0 \times 10^{-3}$ s). Panels (c) and (d) depict the corresponding recovered qGP DRTs, along with their associated 2D contour plots of median DRTs in panels (e) and (f).

remain in handling large-scale datasets efficiently. Future research should explore alternative low-rank approximation techniques beyond SVD to further optimize computational speed and accuracy [63]. Additionally, investigating advanced sampling approaches that enable direct sampling from high-dimensional spaces could potentially accelerate the analysis process [59]. Finally, the integration of scalable GP

approximations such as sparse GPs and distributed GPs, could improve scalability and performance for large-scale or distributed EIS data analysis [36]. These advancements could collectively lead to faster analysis time, greater accuracy, and the ability to handle larger EIS datasets more effectively.

In our current work, we employed evidence maximization for

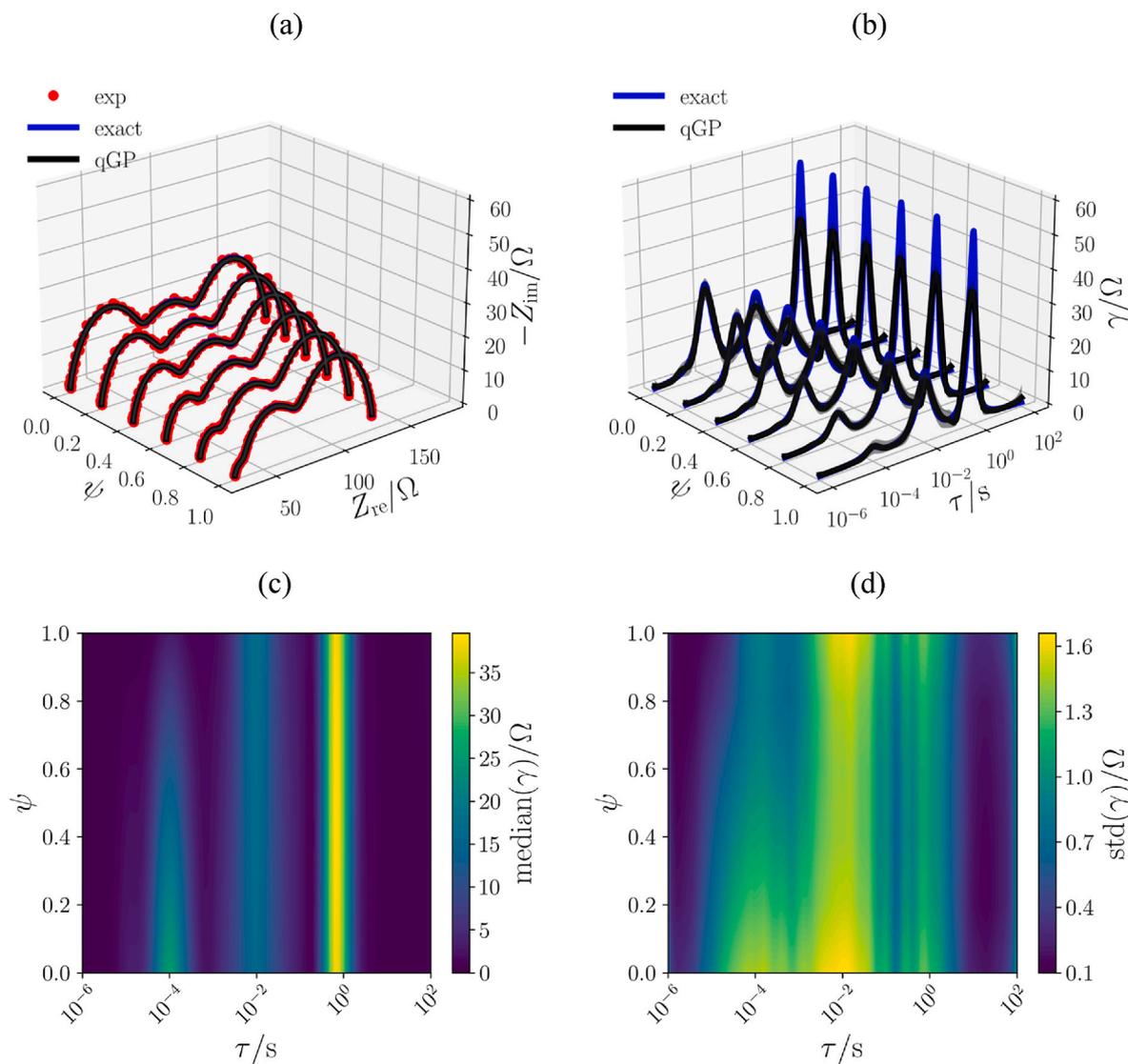


Fig. 5. (a) Nyquist plots of the distant $3 \times \text{ZARC}$ model with $R_{ct,1}(\psi)$ exhibiting a dependence on ψ ($\sigma_n^{\text{exp}} = 0.5\Omega$, $\tau_1 = 1.0 \times 10^{-4} \text{ s}$, $\tau_2 = 1.0 \times 10^{-2} \text{ s}$, and $\tau_3 = 1.0 \times 10^{-1} \text{ s}$). Panel (b) shows the corresponding qGP DRTs, while panels (c) and (d) display 2D contour plots of the median and standard deviation, respectively.

hyperparameter selection [36,64]. However, exploring alternative approaches, such as using products of GP experts as surrogate models, could potentially address scalability challenges, though with a possible trade-off in predictive accuracy [65]. This remains a significant avenue for future research.

A key challenge in our current framework is accurately capturing discontinuities in the DRT, which are essential for representing elements like fractal, Gerischer, and piecewise constant impedance. The inherently smooth nature of GPs poses limitations in this regard. Future research should investigate integrating hierarchical GP methods, such as the deep GPs [66] or other Bayesian methods [55] to enhance the recovery of DRT discontinuities, thereby improving the framework's ability to model complex impedance behaviors and expanding its applicability in EIS data analysis.

5. Conclusions

This work introduces a novel qGP-DRT framework, enabling simultaneous deconvolution of the DRT from EIS data acquired under various experimental conditions. Inheriting key features from the previously developed fGP-DRT model, this approach maintains a probabilistic GP foundation, robustness to experimental noise, and, unlike the vanilla

GP-DRT, full utilization of complete impedance spectra and non-negativity of the recovered DRTs. Notably, the qGP-DRT model simplifies to the fGP-DRT model when considering a single experimental condition. Unlike the fGP-DRT model, the qGP-DRT model can invert the DRT across multiple experimental conditions, effectively bridging dependencies on both frequencies and experimental states. Validation on synthetic and real EIS datasets confirms the consistency and effectiveness of the approach. In short, this article introduces a novel methodology for the probabilistic analysis of multiple EIS spectra, paving the way for further research using GP-based methods for DRT deconvolution.

Code availability

Relevant code is available at <https://github.com/ciuccislab/quaSiGP>.

CRediT authorship contribution statement

Adeleke Maradesa: Writing – review & editing, Writing – original draft, Visualization, Validation, Software, Methodology, Investigation, Formal analysis, Data curation, Conceptualization. **Baptiste Py:** Writing

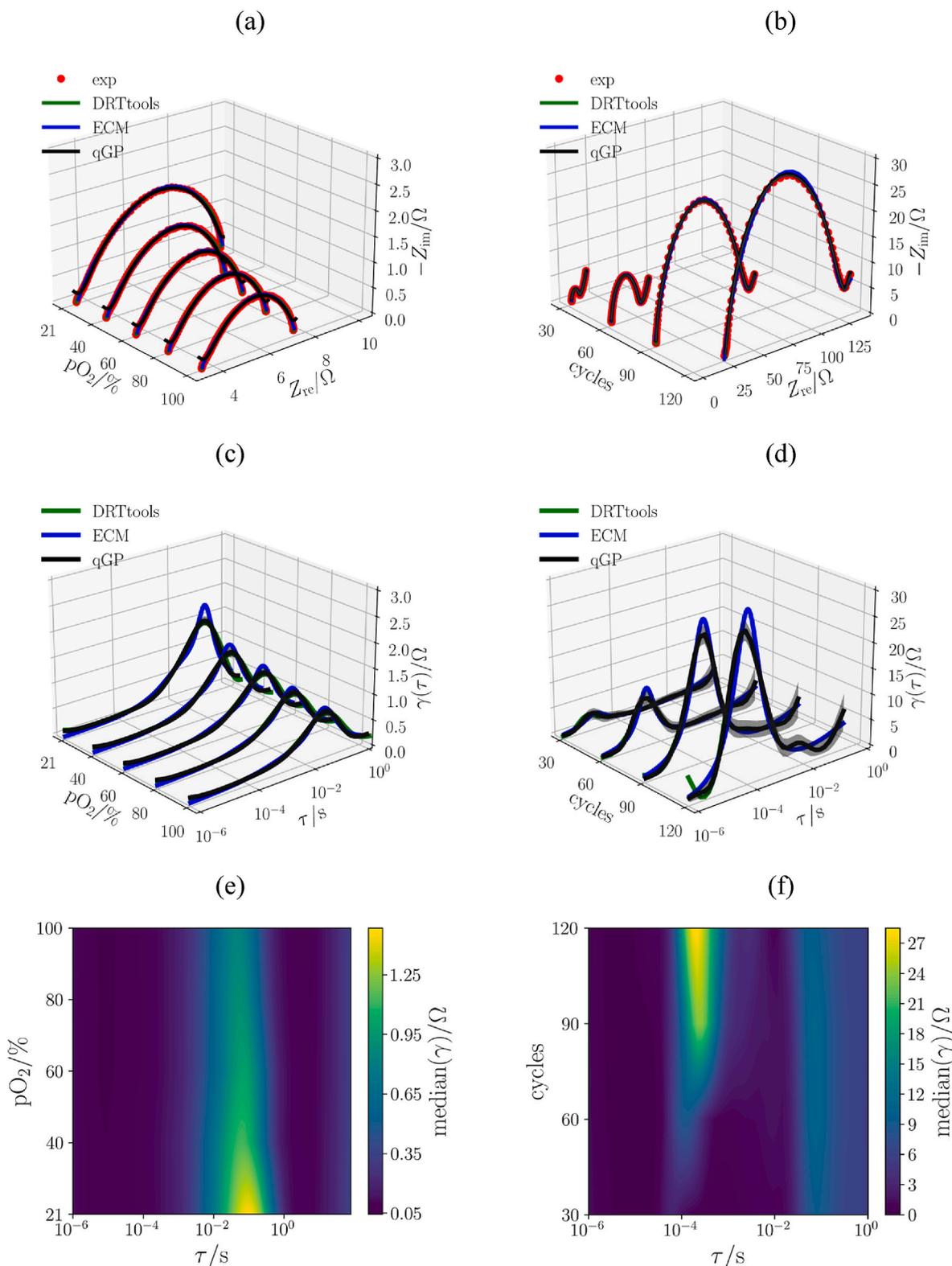


Fig. 6. Nyquist plots of the impedance from fuel cell with BLF (a) and LFP (b) electrodes. Panels (c) and (d) show the corresponding recovered qGP DRTs, along with their associated 2D contour plots for median DRTs in panels (e) and (f).

– review & editing, Methodology. **Francesco Ciucci:** Writing – review & editing, Visualization, Validation, Supervision, Software, Resources, Project administration, Methodology, Investigation, Funding acquisition, Conceptualization.

Declaration of generative AI and AI-assisted technologies in the writing process

To enhance clarity and readability, Gemini Advanced was utilized to assist in the revision process. Following this, the manuscript underwent a comprehensive review and revision by the authors. Dr. Mark Ellwood

further refined the language for clarity and grammatical accuracy. The authors assume full responsibility for the final content of the manuscript.

paper acceptance

Declaration of competing interest

No conflict of interest statement.

Data availability

The companion code will be released on our GitHub channel upon

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Appendix A. Supplementary data

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.jpowsour.2024.235236>.

List of Symbols

Greek letters

Γ	Covariance matrix for all sets of experiments
Γ_{pq}	Covariance matrix for p^{th} and q^{th} experiments
γ	DRT vector
$\gamma(\log \tau, \Psi)$	Distribution of relaxation times
γ_{exact}	Analytical DRT vector
ϵ	Vector of experimental errors
θ	Vector of hyperparameters
μ	Mean vector for the overall experiments
Σ	Covariance matrix for the overall experiments
σ_L	Prefactor for the kernel $k_L(\Psi, \Psi')$
σ_n	GP hyperparameter
σ_n^{exp}	Experimental error
σ_R	Prefactor for the kernel $k_R(\Psi, \Psi')$
σ_r	Retained singular value
σ_v	Prefactor for the kernel $k_v((\log \tau, \Psi), (\log \tau', \Psi'))$
τ	Relaxation time
$\phi_n(\log \tau)$	Basis function
Ψ	Vector of experimental conditions

Latin letters

A	Discretization matrix for the overall experiments
A_p	Discretization matrix for p^{th} experiment
f	Frequency vector
K	Covariance matrix
k	GP kernel
L_0	Inductance
ℓ_f and ℓ_Ψ	Length scales for the kernel $k_v((\log \tau, \Psi), (\log \tau', \Psi'))$
ℓ_L	Length scale for $k_L(\Psi, \Psi')$
ℓ_R	Length scale for $k_R(\Psi, \Psi')$
M	Number of frequencies
M_{tot}	Number of frequencies for all experiments
$\text{MSE}_{\text{norm}, \gamma}$	Normalized DRT mean square error
$\text{MSE}_{\text{norm}, Z}$	Normalized impedance mean square error
N	Number of collocation points
N_Ψ	Number of experimental conditions
R_{ct}	Charge-transfer resistance
R_∞	Ohmic resistance
S_r	The diagonal matrix containing all the singular values
U_r and V_r	The matrices containing the first r columns and rows of the orthogonal matrices U and V
x	Vector of discretized DRTs for the overall experiment
x_p and x_q	Vector of discretized DRTs for p and q experiments
Z	Vector of the recovered impedance
Z_{exact}	Vector of analytical impedance
Z_{exp}	Vector of experimental impedance
Z_p and Z_q	Vector of experimental impedance for p and q experiments

List of Abbreviations

BLF	Ba _{0.95} La _{0.05} FeO _{3-δ}
fGP	finite Gaussian processes
DRT	Distribution of relaxation times
ECM	Equivalent circuit model
EIS	Electrochemical impedance spectroscopy
GPs	Gaussian processes
LFP	LiFePO ₄
pO ₂	Oxygen partial pressure
SoC	State of charge
SVD	Singular value decomposition
qGPs	quasi-GPs
std	Standard deviation

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